

CHEOPS Science Performance Examples of CHEOPS precision using in-flight data

CHEOPS Consortium, 3rd November 2020



Introduction

Data, tools and brief description

- performances
- therefore referred as "Anonymous"
- All the visits were processed with the latest version of the Data Reduction Pipeline (v12.0, see Hoyer et al. 2020)
- Some light curves show instrumental features. Discussing them is beyond the scope of this document
- The light curves shown were obtained using the DEFAULT photometric aperture (circular aperture, radius = 25 px)
- All light curves were analysed with pycheops (V0.9.3) \Rightarrow https://github.com/pmaxted/pycheops
- detected with a signal-to-noise ratio (SNR) of 1
- pycheops
- discrepancies between "real" and "expected" performances

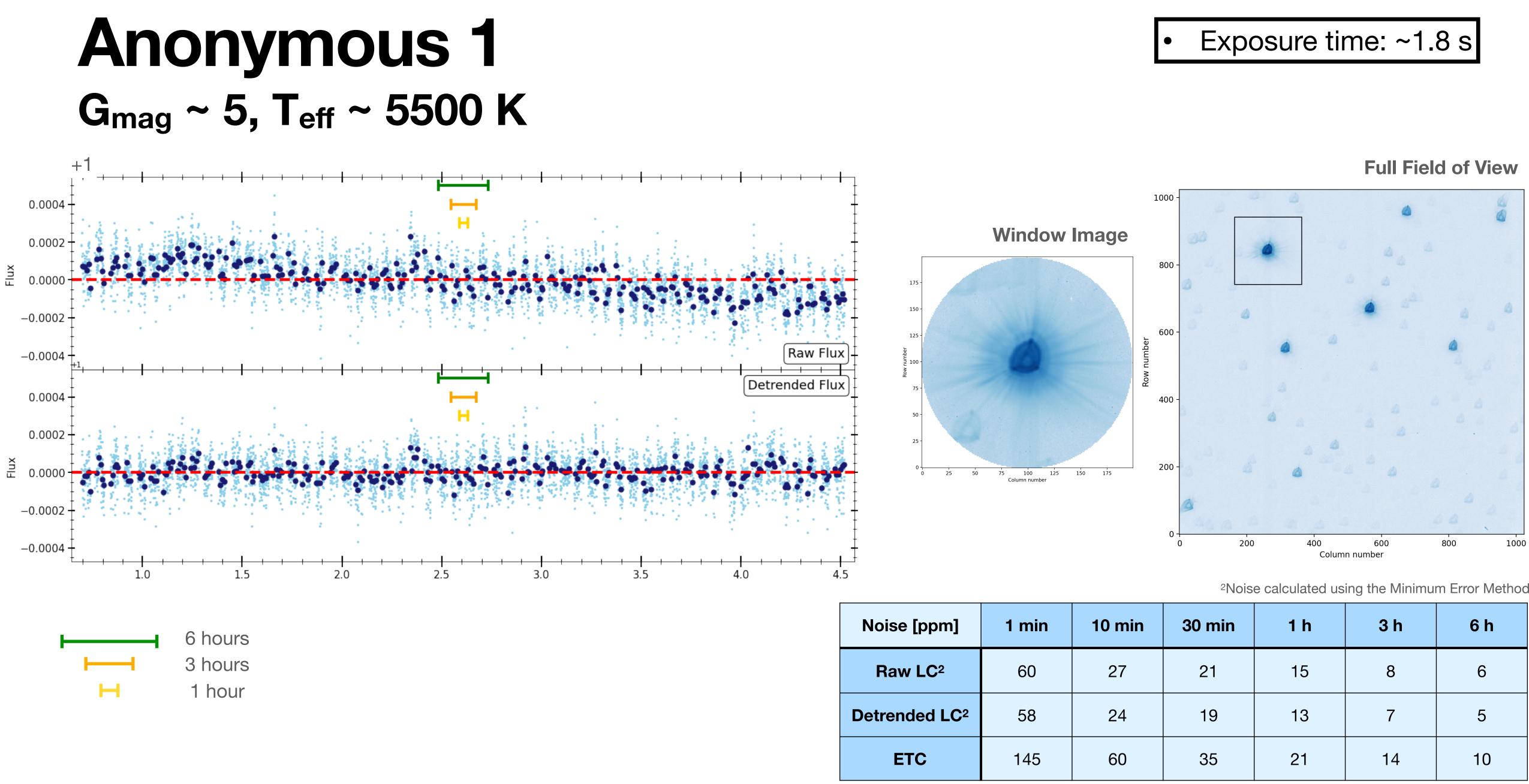
• In the next slides, six visits whose targets cover the nominal CHEOPS magnitude range (6<G_{mag}<12), illustrate CHEOPS

• Three of these targets were observed as part of the GTO programme. For two of them the identities are not provided and

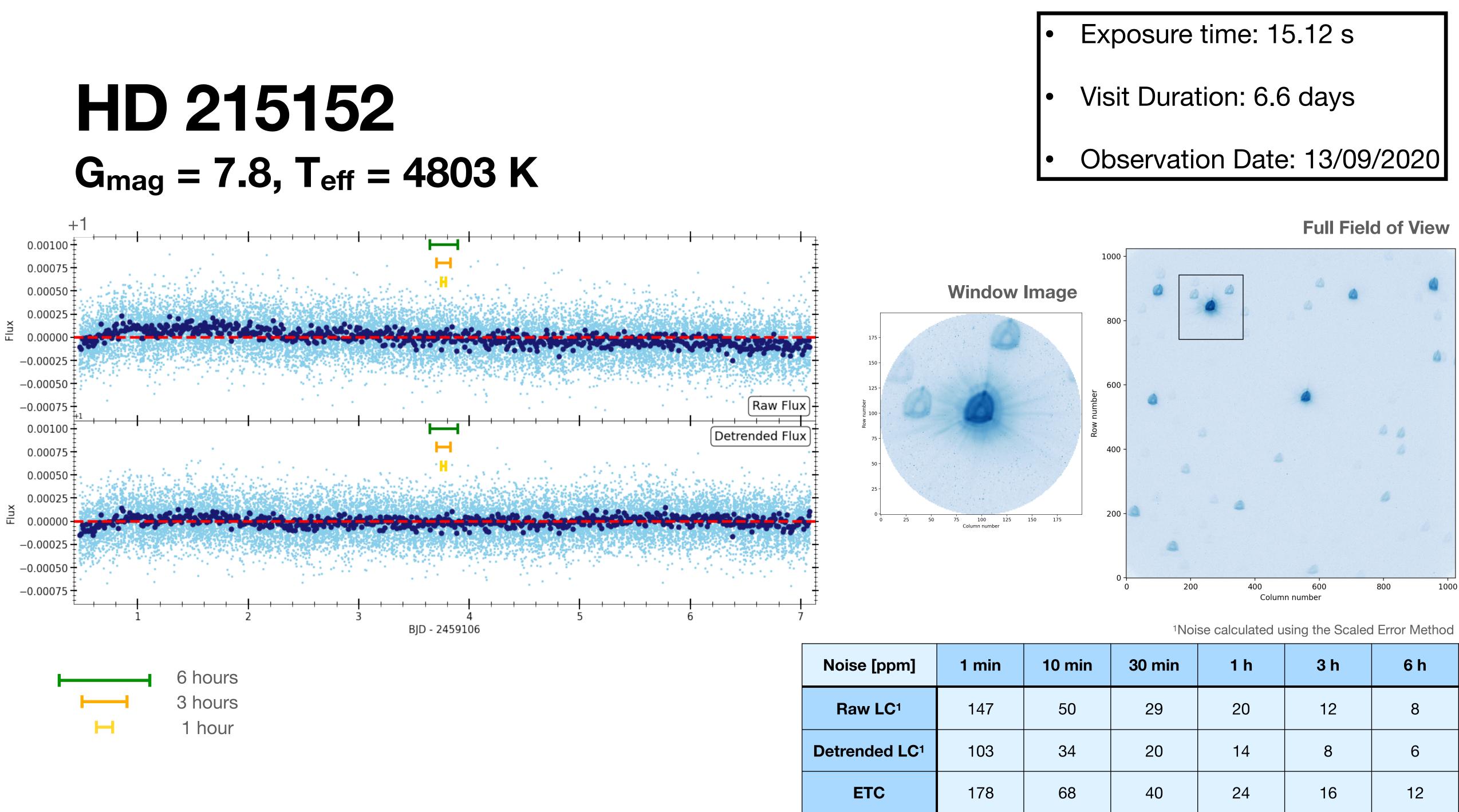
• The science performance is quantified calculating the light curve's noise (in ppm) for six timescales: 1 minute, 10 minutes, 30 minutes, 1 hour, 3 hours and 6 hours The idea behind the noise definition is to determine the transit depth that can be

• Noise estimations are reported for the raw light curve (output of the DRP) and the detrended lightcurve (detrending done using pycheops). The noise estimation corresponds to the Scaled Errors Method, if the scaling factor is lower than 2, or to the Minimum Errors Method otherwise (see the Method description at the end of the slides). Both methods are implemented in

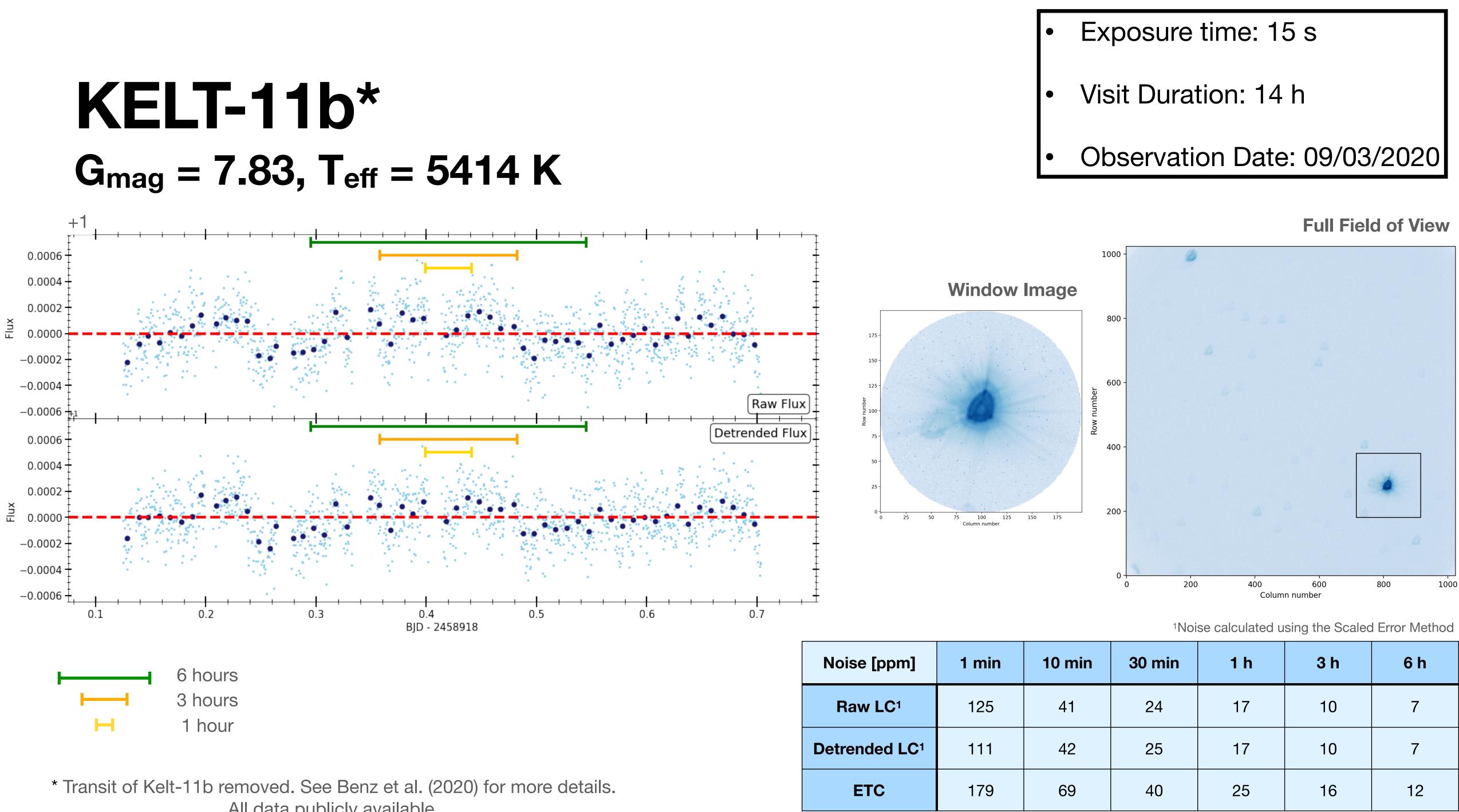
• A comparison with the noise estimation from the ETC (V2.0) is given for reference. Note that for the very bright stars the real performance is in general better than expected, for medium bright stars the ETC agrees well with the measured performance and for faint stars the ETC is reliable only for long time scales (beyond 3 hours). Future versions of the ETC will try to solve the



Noise [ppm]	1 min	10 min	30 min	1 h	3 h	
Raw LC ²	60	27	21	15	8	
Detrended LC ²	58	24	19	13	7	
ETC	145	60	35	21	14	

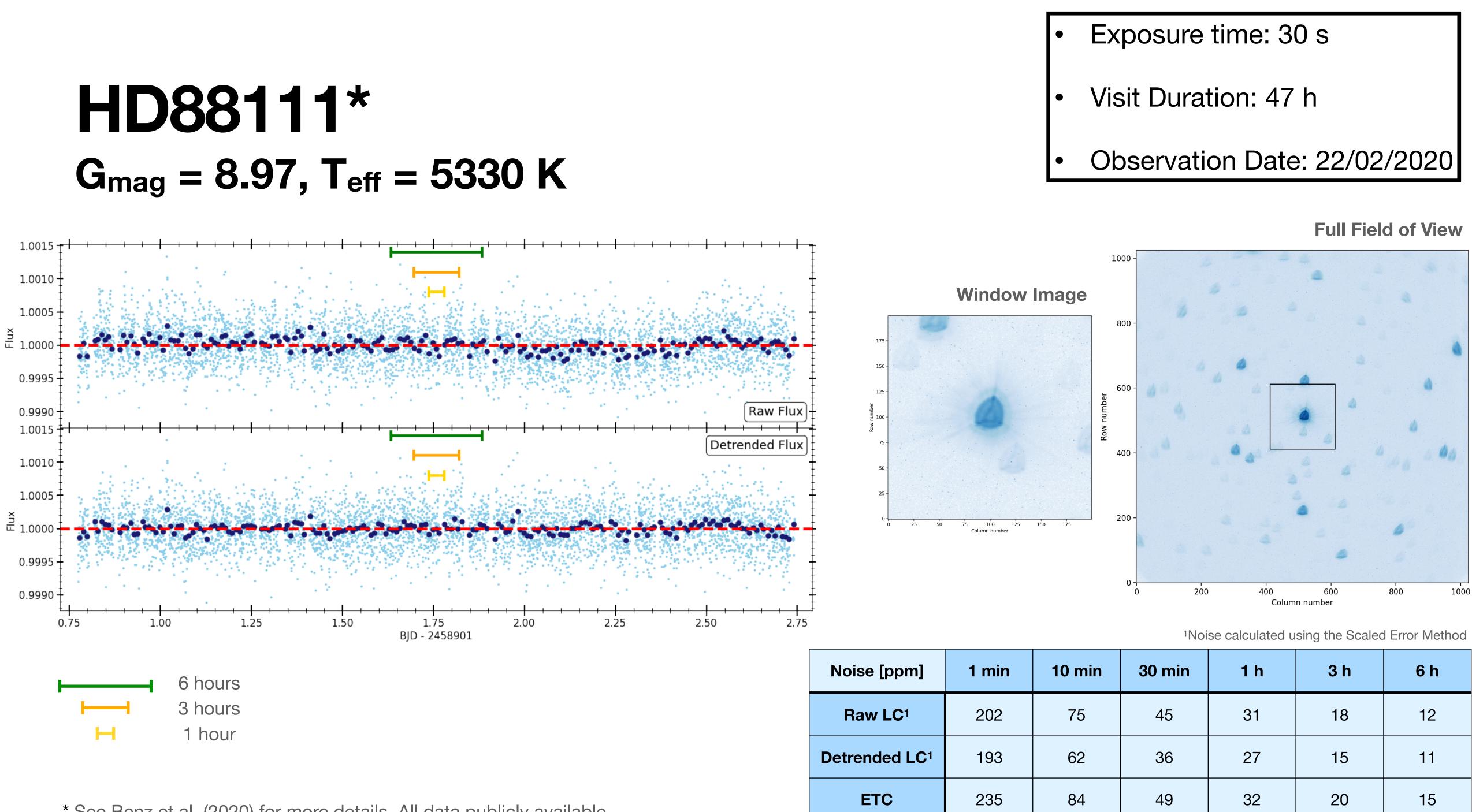


Noise [ppm]	1 min	10 min	30 min	1 h	3 h	
Raw LC ¹	147	50	29	20	12	
Detrended LC ¹	103	34	20	14	8	
ETC	178	68	40	24	16	

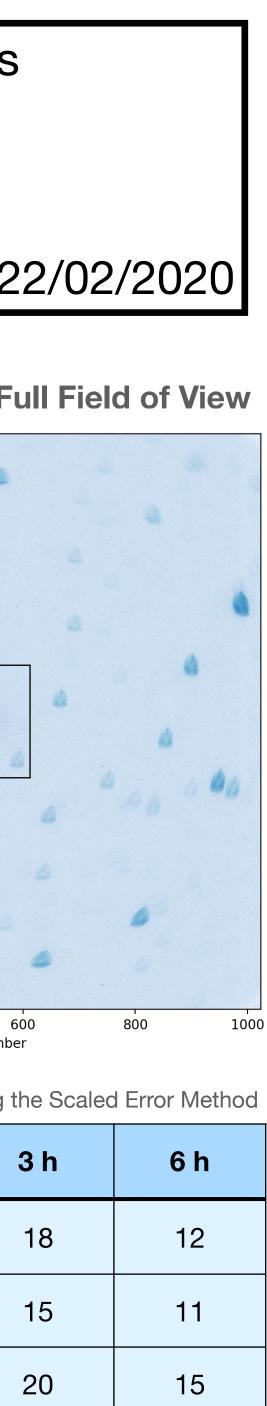


All data publicly available.

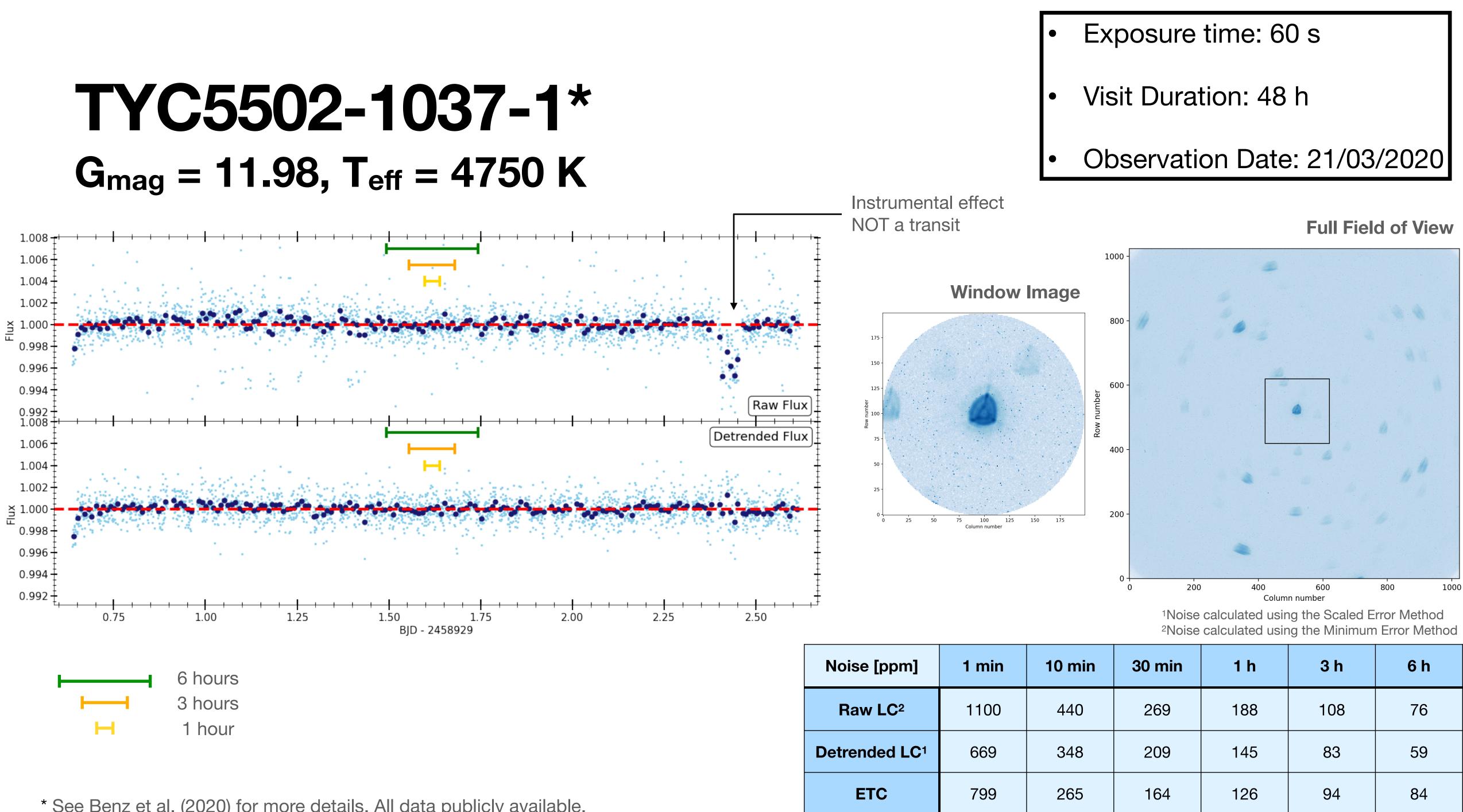
Noise [ppm]	1 min	10 min	30 min	1 h	3 h	
Raw LC ¹	125	41	24	17	10	
Detrended LC ¹	111	42	25	17	10	
ETC	179	69	40	25	16	



* See Benz et al. (2020) for more details. All data publicly available.

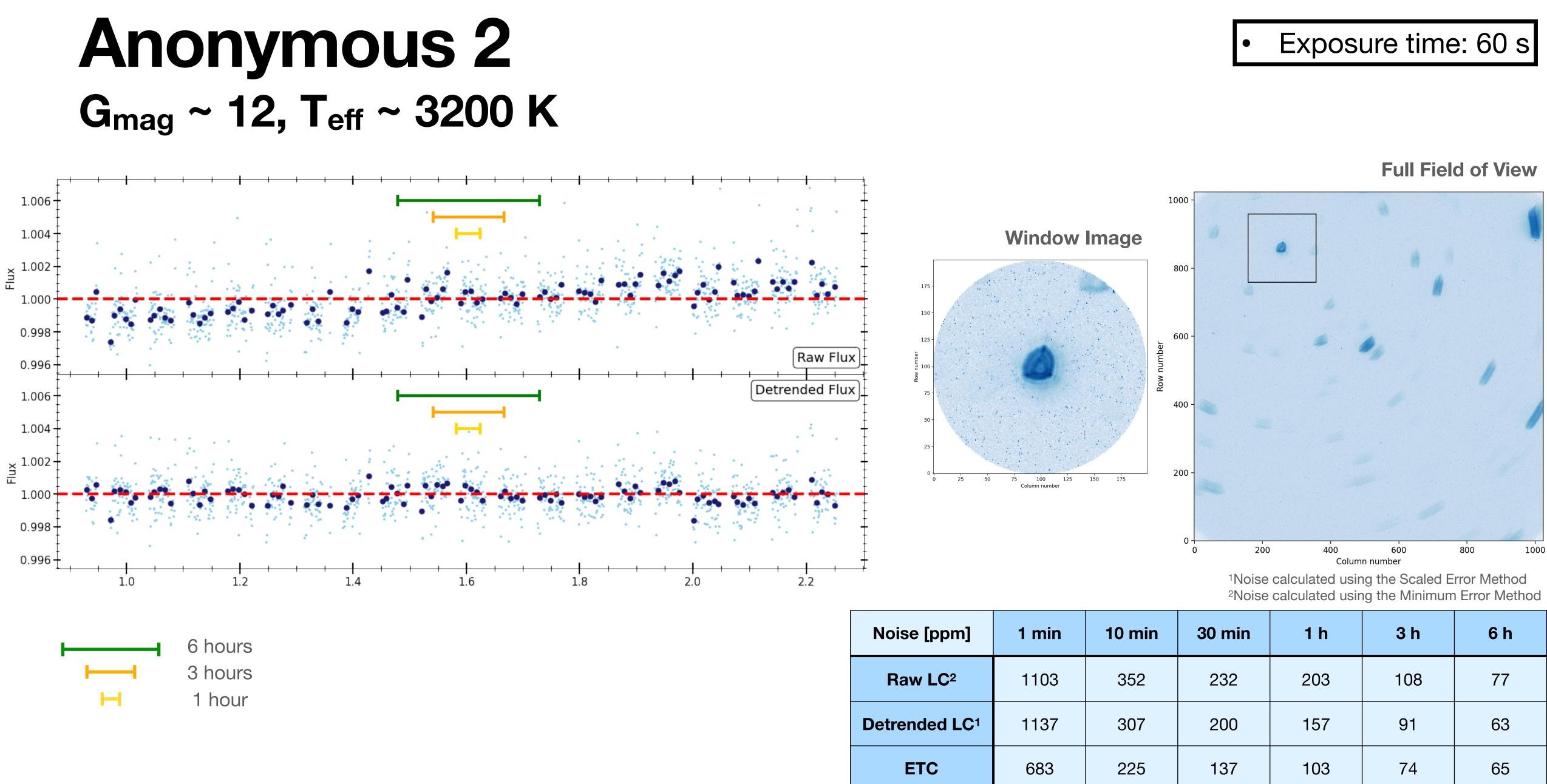


Noise [ppm]	1 min	10 min	30 min	1 h	3 h	
Raw LC ¹	202	75	45	31	18	
Detrended LC ¹	193	62	36	27	15	
ETC	235	84	49	32	20	



* See Benz et al. (2020) for more details. All data publicly available.

Noise [ppm]	1 min	10 min	30 min	1 h	3 h	
Raw LC ²	1100	440	269	188	108	
Detrended LC ¹	669	348	209	145	83	
ETC	799	265	164	126	94	



Noise [ppm]	1 min	10 min	30 min	1 h	3 h	
Raw LC ²	1103	352	232	203	108	
Detrended LC ¹	1137	307	200	157	91	
ETC	683	225	137	103	74	

Noise estimation methods

Method

The photometric precision on the timescale of the transit is estimated by finding the transit depth that can be detected with a signal-to-noise of 1. This is essentially the same method used to calculate the Kepler CDPP noise value.

We can use a factor s to modify the transit depth in a nominal model \mathbf{m}_0 calculated with approximately the correct depth that is scaled as follows ...

$$\mathbf{m}(s) = 1 + s \times (\mathbf{m}_0 - 1).$$

The nominal model is calculated assuming an impact parameter b = 0 and includes limb-darkening appropriate for the star observed.

The data are normalised fluxes $\mathbf{f} = f_1, \ldots, f_N$ with nominal errors $\boldsymbol{\sigma} = \sigma_1, \ldots, \sigma_N$.

Two different methods are used to perform the least-squares fit to the simulated transits that differ in their assumptions concerning the true standard errors on the data.

Scaled errors method The "scaled" method assumes that the actual standard errors are underestimated by some factor b, and that these are normally distributed and uncorrelated, so that the log-likelihood is

$$\ln \mathcal{L} = -\frac{1}{2b^2}\chi^2 - \frac{1}{2}\sum_{k=1}^N \ln \sigma_k^2 - N \ln b - \frac{N}{2}\ln(2\pi),$$

where

$$\chi^2 = \sum_{k}^{N} \frac{(f_k - 1 - s(m_{0,k} - 1))^2}{\sigma_k^2}$$

The maximum likelihood occurs for parameter values \hat{s} , and \hat{b} such that

$$\left. \frac{\partial \ln \mathcal{L}}{\partial s} \right|_{\hat{s}, \hat{b}} = 0$$

and

$$\left. \frac{\partial \ln \mathcal{L}}{\partial b} \right|_{\hat{s}, \hat{b}} = 0$$

from which we obtain

$$\hat{s} = \sum_{k=1}^{N} \frac{(f_k - 1)(m_{0,k} - 1)}{\sigma_k^2} \left[\sum_{k=1}^{N} \frac{(m_{0,k} - 1)^2}{\sigma_k^2} \right]^{-1}$$
$$\hat{b} = \sqrt{\chi^2/N} \quad \text{Scaling factor}$$

and

For the standard errors on these parameters we use

$$\sigma_s^{-2} = -\frac{\partial^2 \ln \mathcal{L}}{\partial^2 s^2} \mid_{\hat{s},\hat{b}}$$

and

$$\sigma_b^{-2} = -\frac{\partial^2 \ln \mathcal{L}}{\partial^2 b^2} \mid_{\hat{s},\hat{b}}$$

$$\sigma_s = b \left[\sum_{i=1}^N \frac{(m_i - 1)^2}{\sigma_i^2} \right]^{-1/2}$$

and

This method is appropriate in cases where the noise scaling factor has a value $b \approx 1$, i.e. unknown sources of noise are small compared to the known sources of noise that are quantied in the flux error values provided.

 $\sigma_b = \left[3\chi^2 / b^4 - N / b^2 \right]^{-1/2}.$

Minimum errors method The "minerr" method is more appropriate where the nominal errors should be treated as a lower-bound to the true standard errors, i.e., where there is some additional noise source that is not well quantified (e.g., poor cosmic-ray rejection). As a result, this method tends to be more conservative, i.e., more pessimisitic. We assume that actual standard error on observation k is σ_k with probability distribution

$$P(\sigma_k | \sigma_{0,k}) = \begin{cases} 0 & \sigma_k < \sigma_{0,k} \\ \frac{\sigma_{0,k}}{\sigma_k^2} & \sigma_k \ge \sigma_{0,k} \end{cases}$$

Assuming independent measurements and uniform priors, the posterior probability distribution is then

$$\ln \mathcal{L} = C + \sum_{k=1}^{N} \ln \left[\frac{1 - \exp(-R_k^2/2)}{R_k^2} \right]$$

where C is a normalising constant and $R_k = (m_k - f_k)/\sigma_{0,k}$ (Sivia & Skilling "Data Analysis", section 8.3.1).

Since this is a function of one parameter only we can be efficiently find the maximum likelihood using a numerical search (Brent's method). The standard error on the s is then found from the value of s that give a log-likelihood that is 0.5 less than the maximum log-likelihood, i.e., at 1-sigma assuming a Gaussian distribution.

